

ДИНАМИЧЕСКИЙ УСЛОВНО-КОРРЕЛЯЦИОННЫЙ АНАЛИЗ ФИНАНСОВОГО РАСПРОСТРАНЕНИЯ: БУДУЩЕЕ ЕВРОДОЛЛАРА И БИТКОИН НА ФОРЭКС-РЫНКАХ

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В данной статье рассматриваются изменяющиеся во времени условные корреляции между будущим рынком евродоллара и семью Форекс-рынками биткоина. Автор применяет двумерную динамическую условную корреляционную модель DCC-GARCH, чтобы зафиксировать потенциальные эффекты финансового распространения между этими рынками в 2017–2019 гг. Эмпирические результаты за исследуемый период показывают финансовое распространение в отношении семи двумерных моделей и потенциальные каналы передачи волатильности между рынками, что имеет решающее значение для директивных органов, которые обеспечивают регулирование вышеуказанных рынков производных финансовых инструментов.

Ключевые слова: модель DCC-GARCH; Форекс-рынки; биткоин; рынок евродоллара; коэффициент финансового распространения; динамически условная коинтеграция.

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DYNAMIC CONDITIONAL CORRELATION ANALYSIS OF FINANCIAL CONTAGION: EURODOLLAR FUTURE AND FOREX BITCOIN MARKETS

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This paper examines the time-varying conditional correlations between eurodollar future market and seven Forex bitcoin markets. We apply a bivariate dynamic conditional correlation DCC-GARCH model in order to capture potential contagion effects between the markets for the period 2017–2019. Empirical results reveal contagion during the under investigation period regarding the seven bivariate model, showing potential volatility transmission channels among the markets. Findings have crucial implications for policymakers, who provide regulations for the above derivative markets.

Keywords: DCC-GARCH model; bitcoin Forex markets; eurodollar future market; financial contagion; dynamic conditional correlations.

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Introduction

This paper investigates the potential volatility spillover and contagion effects of eurodollar future market and seven Forex bitcoin markets. By employing a bivariate DCC-GARCH model, we show significant volatility spillover effects. Moreover, we use the definition of contagion suggested by Forbes and Rigobon¹ according to [1]. Dynamic conditional correlations reveal contagion effects in sub-periods between eurodollar future market and the seven Forex bitcoin markets.

The motivation for this paper is analysed as follows. In the literature, this is the first empirical research exploring the potential conditional second moments of the distribution between eurodollar future market and the seven Forex bitcoin markets [2–4]. Second, we provide new evidence to financial market theory regarding the potential contagion effects between eurodollar future market and the seven Forex bitcoin markets. Third, we use data for Forex bitcoin markets from 2017, when first appeared in datastream database.

There are some empirical studies exploring the conditional volatility of bitcoin market [5–9]. In addition, many researchers support the use of bitcoin market as a speculative market [2; 10–12]. In the literature, there are studies investigating the volatility spillover effects between bitcoin market and different financial markets [7; 9; 13–18]. Moreover, some empirical studies have investigated the effects of future bitcoin markets [11; 19; 20] and the effects between different future and financial markets [21–23]. In this paper, we provide evidence of potential spillovers between markets unexplored before.

The paper is organised as follows. Section two shows the methodology and data. Section three provides the empirical results. The last section of the paper concludes.

Methodology and data

In the first stage, we filter our linear structure of the returns series and decouple of from the conditional variance by employing the VAR model. We generate the daily logarithmic returns.

$$y_t = \gamma + \sum_{s=1}^n \alpha_s r_{t-s} + \varepsilon_t, \text{ with } t=1, \dots, T \text{ and } \varepsilon_t = \sqrt{h_t} u_t,$$

where r_{t-s} is the 2×1 column vector of future and Forex markets, γ and α_s are respectively, a 2×1 vector and 2×2 matrices of parameters and ε_t are 2×1 vectors of innovations. The lag length is chosen by information criteria². u_t is standardised errors, h_t is conditional variance depending on h_t and ε_t for each market lagged one period, generated by the univariate GARCH(1,1) model [24]:

$$h_t = \omega + \alpha \varepsilon_{t-1}^2 + b h_{t-1},$$

where ω is constant, a and b are ARCH and GARCH effects.

Next, we use the R. F. Engle as written in [25] representation of the bivariate GARCH model in order to estimate the bivariate conditional variance matrix (H_t is $N \times N$ matrix, with N the number of markets, $i = 1, \dots, N$) as follows:

$$H_t = D_t R_t D_t$$

where D_t is the conditional variance matrix given by

$$D_t = \begin{pmatrix} \frac{1}{h_{11,t}^2} & \frac{1}{h_{1N,t}^2} \\ \frac{1}{h_{N1,t}^2} & \dots \end{pmatrix},$$

where R_t is the condition correlation matrix of $N \times N$ dimension and defined as follows:

$$R_t = (p_{iit}) = diag\left(q_{11,t}^{-\frac{1}{2}} \dots q_{NN,t}^{-\frac{1}{2}}\right) Q_t diag\left(q_{11,t}^{-\frac{1}{2}} \dots q_{NN,t}^{-\frac{1}{2}}\right),$$

where the $N \times N$ symmetric positive definite matrix $Q_t = (q_{ii,t})$ is given by

$$Q_t = (1 - \alpha - \beta) \bar{Q} + \alpha u_{t-1} u'_{t-1} + \beta Q_{t-1},$$

where \bar{Q} is the $N \times N$ unconditional variance matrix of u_t , and α and β are non-negative scalar parameters, satisfying $\alpha + \beta < 1$.

¹They defined contagion as a significant increase in cross-market linkages after a shock.

²The VAR order length is selected by the final predicted error and the Akaike criterion. The results are available upon request.

We use daily data for eurodollar future market (CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE) and seven Forex bitcoin markets (USD/BITCOIN, EUR/BITCOIN, NZD/BITCOIN, AUD/BITCOIN, CAD/BITCOIN, CHF/BITCOIN, JPY/BITCOIN). We downloaded data from datstream database. We set the period from 18 December 2017 to 20 May 2019 (371 observations) and use the market returns generated by the equation $r_t = \log(p_t) - \log(p_{t-1})$, where p_t is the price of future market on day t and p_{t-1} is the price of future market on day $t-1$.

In table 1 we see the summary statistics for the market returns. CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE exhibits the highest mean value (-6,4221e-005). Based on the highest maximum (0,17818), the minimum (-0,13549) and the highest standard deviation (0,037612) values, EUR/BITCOIN presents the largest fluctuations among all the markets. Additionally, all market returns are negatively skewed, except the case of EUR/BITCOIN. Furthermore, we observe that all market returns show excess kurtosis. In addition, Jarque – Bera statistic results indicate the rejection of the null hypothesis of normality for all market returns except the cases of SGX-KRW/USD CONT.AVG - SETT. PRICE and DGCX-EUR/USD CONTINUOUS AVG. - SETT. PRICE. ADF (Dickey and Fuller 1979) test results reject the null hypotheses of unit root at 1 % level, showing that the daily market returns appropriate for further testing.

Figure 1 graphs the logarithmic returns for CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE, USD/BITCOIN, EUR/BITCOIN, NZD/BITCOIN, AUD/BITCOIN, CAD/BITCOIN, CHF/BITCOIN and JPY/BITCOIN. Based on the virtual observation of the graph, we see time varying levels of fluctuations, indicating the presence of heteroskedasticity and appropriate the use of DCC-GARCH model.

Empirical results

In this section, we present the empirical results generated by the bivariate DCC-GARCH model. Sub-section «Results of the univariate GARCH(1,1) model» shows the results of the univariate GARCH model while in sub-section «Results of the bivariate DCC-GARCH(1,1) model, diagnostic tests and selected information criteria», we analyse the results of the multivariate DCC-GARCH model. In sub-section «Analysis of the dynamic conditional correlations», we report an analysis of the generated dynamic conditional correlations (DCC).

Results of the univariate GARCH(1,1) model. Table 2 shows the estimated values for univariate GARCH(1,1) model. Empirical results report statistically significant ω for CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE and EUR/BITCOIN. Moreover, ARCH (a) and GARCH (b) terms are highly significant for all the markets returns.

In figure 2, we observe the behaviour of conditional variances for CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE, USD/BITCOIN, EUR/BITCOIN, NZD/BITCOIN, AUD/BITCOIN, CAD/BITCOIN, CHF/BITCOIN and JPY/BITCOIN. We see strongly volatile conditional variances for all the market returns over time. Additionally, results indicate a common movement of conditional volatilities.

Results of the bivariate DCC-GARCH(1,1) model, diagnostic tests and selected information criteria. Table 3 presents the results of the bivariate DCC model estimations. the estimated average correlations are statistically significant for the pairs of markets: CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE – NZD/BITCOIN and CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE – CHF/BITCOIN. We see statistically significant β -parameters, indicating strong GARCH effects for the pairs of markets: CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE – USD/BITCOIN, CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE – NZD/BITCOIN and CME-

EURO_GLOBEX_CONT._ - _SETT._ PRICE – CAD/BITCOIN. Additionally, we provide the estimates of the degrees of freedom and of the log-likelihood.

In table 4 we report the results of diagnostic tests and information criteria. $\chi^2(4)$ statistic results suggest that the null hypothesis of no spillovers is rejected at 1 % significance level. Ljung – Box test results [26–27] provide evidence of no serial autocorrelation, suggesting the absence of misspecification errors of the estimated multivariate GARCH model. Moreover, the estimated AIC and SIC information criteria are presented.

Figure 3 plots the conditional covariances for all the pairs of market returns during the whole period. We observe for all the conditional covariances a tremble trend. Additionally, conditional covariances seem to be extreme volatile.

Analysis of the dynamic conditional correlations. Table 4 shows the descriptive statistics of the dynamic conditional correlations of the seven pairs of markets. We observe the highest mean value (0,65938) is for the pair of markets CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE – NZD/BITCOIN. The highest standard deviation value for the pair of markets CME-EURO_GLOBEX_CONT._ - _SETT._ PRICE – CHF/BITCOIN indicates that the specific DCC experiences larger flunctuations. The statistical significant skewness, excess kurtosis and the Jarque – Bera test statistics indicate that the DCCs for all the pairs of markets are not normally distributed.

Table 1

Summary statistics of the daily market logarithmic returns

Criterion	CME-EURO_GLOBEX_CONT - SETT_- PRICE	USD/BITCOIN	EUR/BITCOIN	NZD/BITCOIN	AUD/ BITCOIN	CAD/ BITCOIN	CHF/BITCOIN	JPY/BITCOIN
Mean	-6.4221e-005	-0.0010108	-0.0018578	-0.00091517	-0.00087346	-0.00094496	-0.00096836	-0.0010214
Minimum	-0.0068607	-0.10917	-0.13549	-0.11073	-0.11254	-0.11255	-0.11434	-0.1131
Maximum	0.005153	0.09037	0.17818	0.099148	0.10065	0.099186	0.094779	0.094567
Standard deviation	0.0018299	0.020898	0.037612	0.021073	0.021021	0.020991	0.020999	0.020963
Skewness	-0.13553*	-0.35431**	0.67904***	-0.29452***	-0.28027**	-0.29043***	-0.37228***	-0.35837***
t-Statistic	1.0686	2.7935	5.3539	2.3222	2.2098	2.2899	2.9353	2.8255
p-Value	0.28527	0.0052135	8.6086e-008	0.020223	0.027118	0.022027	0.0033326	0.0047200
Excess Kyr-tosis	0.40483*	3.6739***	3.7132***	3.9233***	4.1678***	4.0652***	4.1225***	4.1409***
t-Statistic	1.6002	14.522	14.677	15.508	16.474	16.068	16.295	16.368
p-Value	0.10956	8.7966e-048	9.0247e-049	3.0827e-054	5.6314e-061	4.2452e-058	1.0695e-059	3.2443e-060
Jarque-Bera	3.6592***	215.83***	241.00***	242.65***	272.64***	259.97***	270.56***	272.27***
p-Value	0.16048	1.3575e-047	4.6570e-053	2.0430e-053	6.2619e-060	3.5285e-057	1.7769e-059	7.5253e-060
ADF Test	-11.2202***	-10.5247***	-9.83784***	-10.3117***	-10.3536***	-10.4758***	-10.3782***	-10.5182***

Note. * – statistical significance at the 10 % level; ** – statistical significance at the 5 % level; *** – statistical significance at the 1 % level.

Source: Datastream® Database.

Table 2

Estimates of univariate GARCH(1,1) model

	CME-EURO_GLOBEX_CONT._SETT._PRICE	USD/BITCOIN	EUR/BITCOIN	NZD/BITCOIN	AUD/BITCOIN	CAD/BITCOIN	CHF/BITCOIN	JPY/BITCOIN
Constant (ω)	1.885387**	0.018679	9.787036***	0.022837	0.015968	0.015335	0.021960	0.023981
t-Statistic	2.288	0.5097	5.215	0.5409	0.4141	0.4090	0.6186	0.6129
p-Value	0.0227	0.6106	0.0000	0.5889	0.6790	0.6828	0.5366	0.5403
ARCH (a)	-0.083052*	0.087080***	0.288417**	0.082228***	0.084434**	0.088373***	0.090383***	0.085605**
t-Statistic	-1.708	3.033	2.703	3.203	2.284	3.303	3.135	2.862
p-Value	0.0885	0.0026	0.0072	0.0015	0.0229	0.0011	0.0019	0.0045
GARCH (b)	0.633162***	0.908741***	-0.043610	0.913403***	0.913877***	0.910357***	0.906721***	0.909705***
t-Statistic	3.422	34.81	-0.7257	39.96	28.30	40.27	37.67	34.52
p-Value	0.0007	0.0000	0.4685	0.0000	0.0000	0.0000	0.0000	0.0000

Note. * – statistical significance at the 10 % level; ** – statistical significance at the 5 % level; *** – statistical significance at the 1 % level.

Source: Datastream® Database.

Table 3

Estimates of the bivariate DCC-GARCH(1,1) model, degrees of freedom, log-likelihood

	CME-EURO_GLOBEX_CONT. SETT_PRICE - USD/ BITCOIN	CME-EURO_GLOBEX_CONT. SETT_PRICE - EUR/ BITCOIN	CME-EURO_GLOBEX_CONT. SETT_PRICE - NZD/ BITCOIN	CME-EURO_GLOBEX_CONT. SETT_PRICE - AUD/ BITCOIN	CME-EURO_GLOBEX_CONT. SETT_PRICE - CAD/ BITCOIN	CME-EURO_GLOBEX_CONT. SETT_PRICE - CHF/ BITCOIN	CME-EURO_GLOBEX_CONT. SETT_PRICE - JPY/BITCOIN
Average COR \bar{ij}	-0.025859	0.003014	0.064134*	0.065718	0.034294	0.056540*	0.027894
t-Statistic	-0.5173	0.06000	1.305	0.8622	0.7022	1.095	0.5174
p-Value	0.6052	0.9522	0.1929	0.3891	0.4830	0.2741	0.6052
alpha (α)	0.0000008	0.0000000	0.0000001	0.091887	0.0000009	0.094385	0.086607
t-Statistic	0.09638	0.00	0.1350	0.1179	0.04310	0.5050	0.2195
p-Value	0.9233	1.0000	0.8927	0.9062	0.9656	0.6138	0.8264
beta (β)	0.829406***	0.237771	0.640912*	0.000000	0.671466**	0.000000	0.000000
t-Statistic	3.008	0.4349	1.679	0.00	2.058	0.00	0.00
p-Value	0.0028	0.6639	0.0941	1.0000	0.0403	1.0000	1.0000
Degrees of freedom (df)	4.233330***	4.874488***	4.382709***	4.235011***	4.308371***	4.262997***	4.353704***
t-Statistic	6.301	5.294	6.229	6.429	6.392	6.318	6.313
p-Value	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
Log-likelihood	2772.508	2535.800	2765.063	2768.983	2768.140	2770.719	2768.663

Note. * – statistical significance at the 10 % level; ** – statistical significance at the 5 % level; *** – statistical significance at the 1 % level.

Source: Datastream® Database.

Table 4

Diagnostic tests and information criteria

	CME-EU-RO_GLOBEX_-CONT._-SETT._PRICE - USD/ BITCOIN	CME-EU-GLOBEX_CONT._-SETT._PRICE - EUR/ BITCOIN	CME-EU-RO_GLOBEX_-CONT._-SETT._PRICE - NZD/ BITCOIN	CME-EU-RO_GLOBEX_-CONT._-SETT._PRICE - CAD/ BITCOIN	CME-EU-RO_GLOBEX_-CONT._-SETT._PRICE - CHF/ BITCOIN	CME-EU-RO_GLOBEX_-CONT._-SETT._PRICE - JPY/ BITCOIN
χ^2	584.62**	103.74**	442.70**	547.57**	631.75**	519.70**
p-Value	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
Hosking (50)	187.941	193.790	195.408	197.594	193.698	196.431
p-Value	0.7197109	0.6103895	0.5784710	0.5348331	0.6121863	0.5580847
Hosking ² (50)	155.988	142.113	156.666	151.207	152.251	151.778
p-Value	0.9877142	0.9990083	0.9863842	0.9943476	0.9932513	0.9937679
Li-McLeod (50)	189.902	195.818	197.108	199.168	195.358	198.127
p-Value	0.6844528	0.5703171	0.5445508	0.5033225	0.5794607	0.5241589
Li-McLeod ² (50)	158.881	146.652	159.689	154.424	155.655	155.130
p-Value	0.9811695	0.9975265	0.9788984	0.9903752	0.9883283	0.9892417
Akaike	0.013550	0.017008	0.013659	0.013601	0.013614	0.013576
Schwarz	0.119320	0.122778	0.119429	0.119372	0.119384	0.119346

Note. ** – statistical significance at the 5 % level.

Source: Datastream® Database.

Table 5

Statistical properties of the Multivariate GARCH-DCC's

Criterion	CME-EURO – GLOBEX – CONT. – SETT. – PRICE – USD/ BITCOIN	CME-EURO – GLOBEX – CONT. – SETT. – PRICE – EUR/ BITCOIN	CME-EURO – GLOBEX – CONT. – SETT. – PRICE – NZD/ BITCOIN	CME-EURO – GLOBEX – CONT. – SETT. – PRICE – CAD/BITCOIN	CME-EURO – GLOBEX – CONT. – SETT. – PRICE – CHF/ BITCOIN	CME-EURO – GLOBEX – CONT. – SETT. – JPY/ BITCOIN
Mean	-0.025859	0.0030141	0.064134	0.063136	0.034294	0.025399
Minimum	-0.025859	0.0030141	0.064134	0.29711	0.034294	-0.31205
Maximum	-0.025859	0.0030141	0.064134	0.28725	0.034294	0.28576
Standard deviation	1.3411e-009	2.8696e-019	7.4978e-018	0.067811	4.9471e-018	0.069185
Skewness	-1.4290***	-0.11195	0.00000	-0.83641***	-2.1852***	-0.88398***
p-Value	1.9060e-029	0.37743	1.0000	4.2601e-011	1.5946e-066	3.1739e-012
Excess kurtosis	6.9849***	-0.37769*	0.42593*	6.0377***	2.5339***	6.1992***
p-Value	8.5871e-168	0.13547	0.092267	7.0067e-126	1.3008e-023	1.3583e-132
Jarque – Bera	878.10***	2.9719	2.7968	605.15***	393.46***	640.64***
p-Value	2.1051e-191	0.22628	0.24699	3.9282e-132	3.6477e-086	7.6923e-140
						1.9583e-166

Note. * – statistical significance at the 10 % level; *** – statistical significance at the 1 % level.

Source: Datastream® Database.

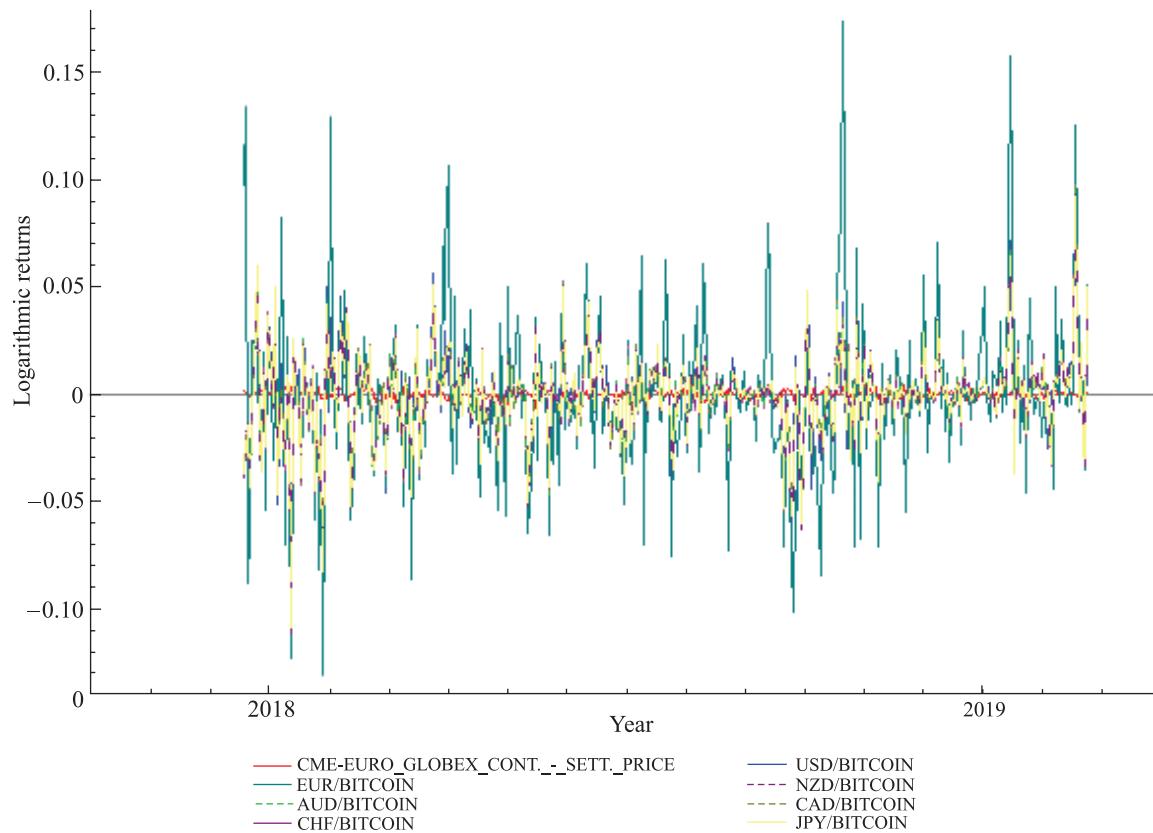


Fig. 1. Actual series of the logarithmic returns of the markets

Source: Datastream® Database

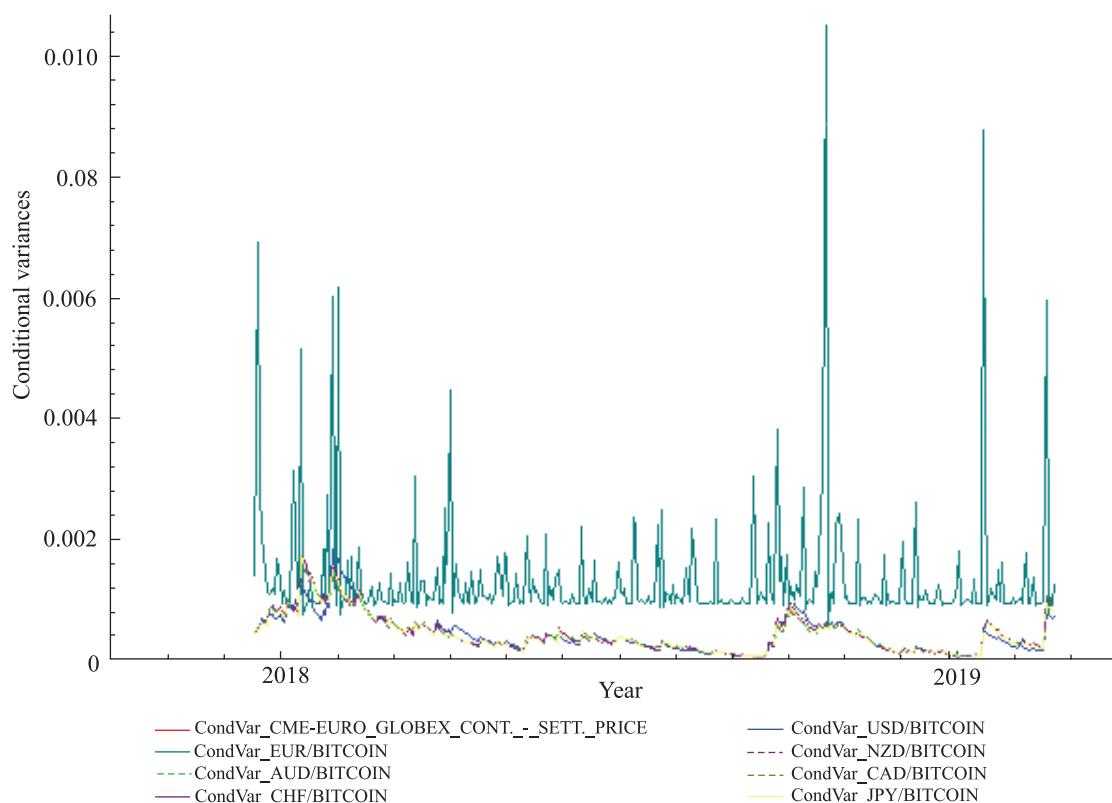


Fig. 2. Conditional variances of the univariate GARCH(1,1) model

Source: Datastream® Database

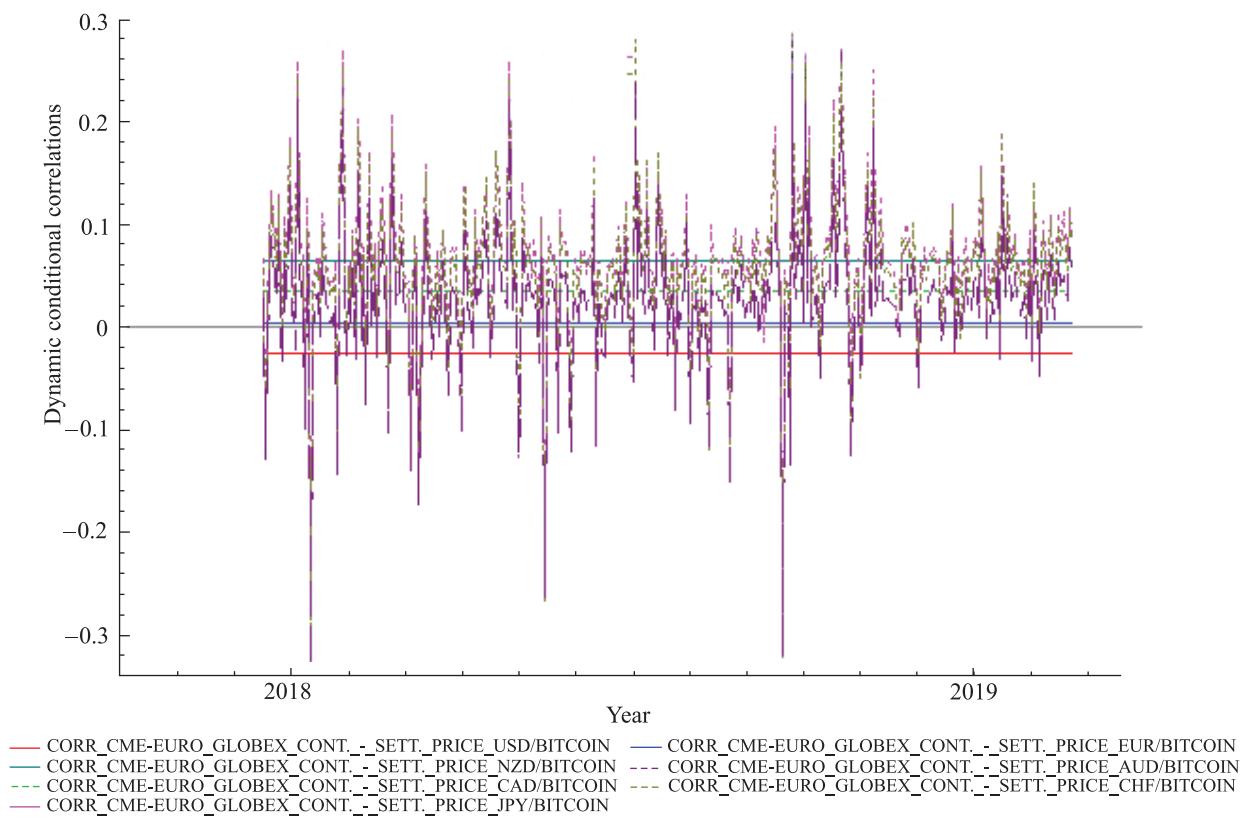


Fig. 3. Conditional covariances of the bivariate DCC-GARCH(1,1) model

Source: Datastream® Database

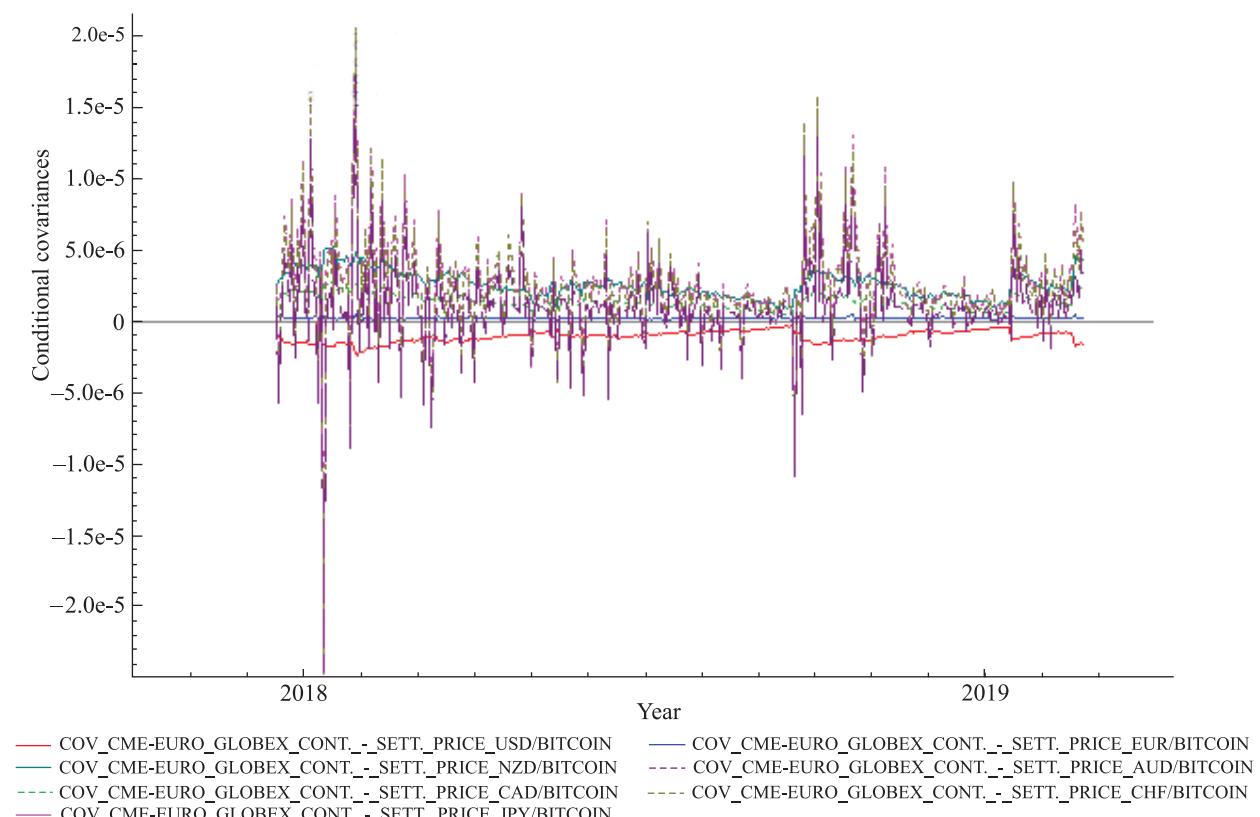


Fig. 4. Dynamic conditional correlations of the bivariate DCC-GARCH(1,1) model

Source: Datastream® Database

Figure 4 shows the pair-wise DCCs. DCCs present positive values in sub-periods, supporting the existence of contagion, implying the specific correlations risky for any investor. Additionally, we can notice the effects of major economic events on the DCC graphs as we see that the lines are bouncing above and beyond and are extreme volatile for some pairs of markets.

Conclusions

This paper investigates the potential volatility spillovers effects and the existence of contagion effects among eurodollar future market and seven Forex bitcoin markets by employing a bivariate DCC-GARCH model. We set the under investigation period from 2017 until 2019. To the best of our knowledge, this is the first empirical study, investigating volatility spillovers between eurodollar future market and five Forex bitcoin markets.

The main empirical results are summarised as follows. Based on the descriptive statistics, EUR/BITCOIN returns presents the largest fluctuations compared to the rest markets. Furthermore, results of bivariate DCC-GARCH models indicate strong evidence of volatility spillover effects. DCCs analysis shows evidence of strong co-movements for some pairs of markets. Additionally, DCCs reveal contagion for some pairs of markets in sub-periods. The empirical results are of interest to policymakers, who provide regulations for the under investigation derivative markets as well as to market-makers.

References

1. Forbes K, Rigobon R. No contagion, only interdependence: measuring stock market comovements. *The Journal of Finance*. 2002;57(5):2223–2261. DOI: 10.1111/0022-1082.00494.
2. Aalborg H, Aarhus PM, de Vries JE. What can explain the price, volatility and trading volume of bitcoin? *Finance Research Letters*. 2019;29:255–265. DOI: 10.1016/j.frl.2018.08.010.
3. Ardia D, Bluteau K, Rüede M. Regime changes in bitcoin garch volatility dynamics. *Finance Research Letters*. 2019;29: 266–271. DOI: 10.2139/ssrn.3180830.
4. Bouoiyour J, Selmi R. Bitcoin: A beginning of a new phase. *Economics Bulletin*. 2016;36(3):1430–1440.
5. Bouri E, Molar P, Azzi G, Rouband D, Hagfors LI. On the hedge and safe haven properties of bitcoin: is it really more than a diversifier? *Finance Research Letters*. 2017;20:192–198. DOI: 10.1016/j.frl.2016.09.025.
6. Charles A, Olivier D. Volatility estimation for bitcoin: replication and robustness. *International Economics*. 2019;157:23–32. DOI: 10.1016/j.inteco.2018.06.004.
7. Dyhrberg AH. Bitcoin, gold and the dollar – a garch volatility analysis. *Finance Research Letters*. 2016;16:85–92. DOI: 10.1016/j.frl.2015.10.008.
8. Katsiampa P. Volatility estimation for bitcoin: a comparison of garch models. *Economics Letters*. 2017;158:3–6. DOI: 10.1016/j.econlet.2017.06.023.
9. Wu Sh, Tong M, Yang A, Derbali A. Does gold or bitcoin hedge economic policy uncertainty? *Finance Research Letters*. 2019;31:171–178. DOI: 10.1016/j.frl.2019.04.001.
10. Bouoiyour J, Selmi R. What does bitcoin look like? *Annals of Economics and Finance*. 2015;16(2):449–492.
11. Corbet S, Lucey B, Peat M, Vigne S. Bitcoin futures – what use are they? *Economics Letters*. 2018;172:23–27. DOI: 10.1016/j.econlet.2018.07.031.
12. Yermack D. Is bitcoin a real currency? An economic appraisal. In: David Lee Kuo Chuen, editor. *Handbook of Digital Currency*. New York: University Stern School of Business; National Bureau of Economic Research, 2015. p. 31–43.
13. Dastgir S, Demir E, Downing G, Gozgor G, Marco Lau ChK. The causal relationship between bitcoin attention and bitcoin returns: Evidence from the copula-based granger causality test. *Finance Research Letters*. 2019;28:160–164. DOI: 10.1016/j.frl.2018.04.019.
14. Kristoufek L. Bitcoin meets google trends and Wikipedia: quantifying the relationship between phenomena of the internet era. *Scientific Reports*. 2013;3(1):1–7. DOI: 10.1038/srep03415.
15. Panagiotidis T, Thanasis S, Orestis V. The effects of markets, uncertainty and search intensity on bitcoin returns. *International Review of Financial Analysis*. 2019;63:220–242. DOI: 10.1016/j.irfa.2018.11.002.
16. Dyhrberg AH. Hedging capabilities of bitcoin. Is it the virtual gold? *Finance Research Letters*. 2016;16:139–144. DOI: 10.1016/j.frl.2015.10.025.
17. Bouri E, Gupta R, Tiwari AK, Roubaud D. Does bitcoin hedge global uncertainty? Evidence from wavelet-based quantile-in-quantile regressions. *Finance Research Letters*. 2017;23:87–95. DOI: 10.1016/j.frl.2017.02.009.
18. Demir E, Gozgor G, Marco Lau CK, Vigne SA. Does economic policy uncertainty predict the bitcoin returns? An empirical investigation. *Finance Research Letters*. 2018;26:145–149. DOI: 10.1016/j.frl.2018.01.005.
19. Ruozhou L, Wan Sh, Zili Zh, Xuejun Zh. Is the introduction of futures responsible for the crash of bitcoin? *Finance Research Letters*. 2020;34:101259. DOI: 10.1016/j.frl.2019.08.007.
20. Kim W, Lee J, Kang K. The effects of the introduction of bitcoin futures on the volatility of bitcoin returns. *Finance Research Letters*. 2020;33:101204. DOI: 10.1016/j.frl.2019.06.002.
21. Sebastião H, Godinho P. Bitcoin futures: an effective tool for hedging cryptocurrencies. *Finance Research Letters*. 2020; 33:101230. DOI: 10.1016/j.frl.2019.07.003.
22. Tsiaras K. Dynamic relationship between future FOREX markets in the post Global Financial Crisis. *Journal of Quantitative Methods*. 2020;4(1):30–52. DOI: 10.29145/2020/jqm/040102.
23. Tsiaras K. Volatility spillover and contagion effects between eurodollar future and zero coupons markets: evidence from Italy. *The European Journal of Applied Economics*. 2020;17(2):67–88. DOI: 10.5937/ejae17-26893.

24. Bollerslev T. Generalized autoregressive conditional heteroskedasticity. *Journal of Econometrics*. 1986;31(3):307–327. DOI: 10.1016/0304-4076(86)90063-1.
25. Engle RF. Dynamic conditional correlation-a simple class of multivariate GARCH models. *Journal of Business & Economic Statistics*. 2002;20:339–10.2307/2287656350. DOI: 10.1198/073500102288618487.
26. Hosking JRM. The Multivariate Portmanteau Statistic. *Journal of the American Statistical Association*. 1980;75(371):602–608. DOI: 10.2307/2287656.
27. McLeod AI, Li WK. Diagnostic checking ARMA time series models using squared-residuals autocorrelations. *Journal of Time Series Analysis*. 1983;4(4):269–273. DOI: 10.1111/j.1467-9892.1983.tb00373.x.

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